# UNITED STATES DEPARTMENT OF THE INTERIOR GEOLOGICAL SURVEY 

# Algorithm and BASIC Program for Ordinary Least Squares Regression in Two and Three Dimensions by Gary R. Olhoeft 

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Equations must frequently be fitted to data to make the data more easily usable (Bevington, 1969; Searle, 1971; Daniel and Wood, 1971). Most least squares fitting algorithms are written for equations in two dimensions. A typical linear example is
$y=f(x)=b_{0}+b_{1} x+b_{2} x^{2}+\ldots+b_{k} x^{k}$
where the data are in pairs of $x, y$ points and the coefficients, $b_{i}$, are to be found. This type of equation, which is linear in the coefficients, may be represented in matrix form as
$\gamma=X B$
where all capital letters are matrices, and
$Y$ is $n$ by 1 in size ( $n$ data points).
$X$ is $n$ by $k+1$ in size ( $k+1$ coefficients in $b$ )
and $B$ is $k+1$ by 1 .
The desired solution to the above equation is the coefficient matrix $B$. The expectation value of the coefficient matrix is

$$
\underline{B}=\left(X^{\top} X\right)^{-1} X^{\top} Y
$$

It must be noted that the inversion of $X^{\top} X$ may be 111 conditioned for large $k$ unless the arithmetic precision of the computer is sufficiently high. The programs illustrated here require 48-bit arithmetic. The development of this type of ordinary least squares fit is adequately derived in Searle (1971). Table 1 summarizes the resultant findings. Program 1 illustrates a typical use: fitting electrical conductivity data of aqueous solutions to their concentration dependence using the equation

$$
y=b_{0}+b_{1} x^{1 / 2}+b_{2} x \ln x+b_{3} x+b_{4} x^{3 / 2}
$$

However, it is sometimes desirable to fit a three-dimensional equation, $y=f(x, z)$, such as the electrical conductivity dependence
upon both concentration and temperature. In this case, the matrix formulation becomes

$$
Y=X B Z
$$

with the expectation

$$
\underline{B}=\left(X^{\top} X\right)^{-1} X^{\top} Y Z^{\top}\left(Z Z^{\top}\right)^{-1} \text {. }
$$

The extension of the two dimensional model to three dimensions is fairly straightforward using Searle (1971). The coefficients in the $B$ matrix must still be linear in both $X$ and $Z$, and the matrices may become ill conditioned if insufficient precision is used. Program 2 illustrates the three-dimensional example.

Both programs 1 and 2 are written in Hewlett-Packard 9830 BASIC*, but may be easily adapted to any language that has the appropriate matrix functions and precision.

## REFERENCES:

Bevington, P. R., 1969, Data Reduction and Error Analysis for the
Physical Sciences: New York, McGraw-Hill, 336 p.
Daniel, C. and Wood, F. S., 1971, Fitting Equations to Data:
New York, Wiley-Interscience, 342 p.
Searle, S. R., 1971, Linear Models: New York, Wiley, 532 p.

[^0]TABLE 1 : Ordinary least squares regression (intercept model, Searle, 1971) $y(x)=b_{0}+b_{1} x+\ldots+b_{k} x^{k}$ (or other equation linear in coefficients $b$ ) $Y=X B$ where $Y$ is $n$ by $1, X$ is $n$ by $k+1$, and $B$ is $k+1$ by 1 $\underline{B}=\left(X^{\top} X\right)^{-1} X^{\top} Y=$ expectation value of coefficient matrix $B$ $\operatorname{var}(\underline{B})=\left(x^{\top} x\right)^{-1} \sigma^{2}=$ variance in expectation value of $B$ $\sigma^{2}=$ square of standard deviation of original data $\underline{Y}=X \underline{B}=$ estimated expectation value of $Y$
$\operatorname{var}(\underline{Y})=x\left(x^{\top} \dot{X}\right)^{-1} x^{\top} \sigma^{2}$
SSE $=(Y-\underline{Y})^{\top}(Y-\underline{Y})=$ residual error sum of squares $=Y^{\top} Y-\underline{B}^{\top} X^{\top} Y$
$\sigma^{2}=$ SSE/(n-k-1) $=$ square of expectation value of standard deviation for $n$ data points and $k+1$ coefficients
SST $=Y^{\top} Y=$ total sum of squares
SSR $=$ SST - SSE $=\underline{B}^{T} X^{\top}{ }^{T}=$ sum of squares due to regression
$S S M=n \bar{y}^{2}=$ sum of squares of mean
$S S R_{m}=S S R-S S M=$ regression sum of squares corrected for the mean
$S S T_{m}=S S T-S S M=$ corrected sum of squares
$R^{2}=$ SSR/SST = square of multiple correlation coefficient for $b_{0}=0$
$=$ fraction of total sum of squares accounted for by fitting the model
$R_{m}^{2}=S S R_{m} S S T_{m}=R^{2}$ for $b_{o} \neq 0$

```
Program 1 : Two-dimensional ordinary least squares regression
Lines 10-20 Dimension arrays for equation of form \(Y=X A\)
    30-40 Input a title for the printout
    50-100 Enter the data in \(x, y\) pairs: this program has been
        abbreviated to illustrate the principle. As shown, it
        requires exactly \(30 x, y\) pairs to be entered, but it may
        be easily modified to accept an arbitrary number.
    110-150 Solution for coefficient matrix, A.
    160-260 Solution for SSR (S3), SST (S2), SSE (S1), standard
        deviation (V2) and \(R * R(R 2)\).
    270-400 Printout coefficient matrix, raw data, expectation values
        of \(y\), and differences.
    410-480 Function \(y(x)\) to be fitted.
```

NOTES: Sufficient arithmetic precision must be used to prevent $X^{\top} X$ from becoming ill zonditioned: 48-bit is recommended. To alter the program for use with a different number of $x, y$ pairs, change the dimensions that are equal to 30 in lines $10,20,50$, and 360 . For an arbitrary number, place a test at line 71 to exit the FOR-NEXT loop at an agreed value of $x, y$. Set the $X$ and $Y$ matrices equal to zero at line 21 , and set the dimension that is 30 in line 360 equal to the variable I.
To change the function, alter lines 420-460. If more or less than 5 parameters in $x$ are desired, change the dimension that is 5 in lines $10,20,250,290,300$, and 330 as appropriate.

Program 1 : Two-dimensional ordinary least squares regression


Program 2 : Three-dimensional ordinary least squares regression
lines 10-20 Dimension arrays for equation of the form $S=C B T$.
30-50 Input a title for the printout.
60-230 Input the data: for 5 concentrations, 16 pairs of temperatures and conductivites are entered.
240-330 Solve for coefficient matrix B.
340-350 Solve for estimated expectation values of $S$ (given in the 0 matrix).
360-550 Print the results: raw data, estimated data, and the difference in percent as well as the coefficient matrix, 8.

NOTES: Again; for brevity and to demonstrate the principle, much of the program has been reduced to essentials. This program, as shown, requires exactly 5 concentrations and 16 conductivity-temperature pairs at each concentration. These may be changed by appropriately modifying lines $10,20,60,140$, and $350-480$. The functions to be fit for $C$ are in lines $90-130$ and for $T$ in lines $170-210$. These may be rewritten as functions as shown in Program 1. Both the conductivity dependence upon concentration and temperature are shown with 5 variables in this listing. If that is altered to be more or less than 5 , then lines $10,20,360-480$, and $510-540$ must also be altered.

Program 2 : Three-dimensional ordinary least squares regression.

```
10 DIA E [5,5],T$[8C],N[5,5],C[5,16],C[5,16]
```



```
30 UISE "gRIMTUUT 'IITLE";
40 INFCT TS
50 PRINT T$
60 ROR I=1 TU 5
70 DISP "CCivCLH'IRAIION";
80 INUUL C
90C[I,1]=C
100C[ [, 2]=C*SQRC
11U C[1,3]=C^2*LOCC
120C[ [I,4]=C^2
130C[I,5]=C*C^(3/2)
140 FOR J=1 TO 16
150 EISF C;" TEMFERATUFE, CONDCCTIVITY ";
160 INEUT T,S[I,J]
170 T[1,J]=1
180 T[2,J]=1/T
190 T[3,J]=T
200 T[4,J]=T^2
210 T[5,コ]=T^3
220 NEXT J
230 IWEXT I
240 !!AT E=TRN(C)
250 NAT F=TRA:(T)
260 i:AT B=E*C
270 NAT G=INV (B)
280 HAT N=E*E
290 MAT C=N*S
300 MAT H=&*F
310 MAT E=T*F
320 \AT G=INV(B)
330 NAT E=H*G
340 [AT Li=C*E
350 MAT 0=D*'T
360 WRITE (15,450)0,C[1,1],C[2,1],C[3,1],C[4,1],C[5,1]
370 وRINT
380 FOR J=1 TO 16
390 WRITE (15,450)T[3,J],S[1,J],S[2,J],S[3,J],S[4,J],S[5,J]
400 KRITE (15,450)T[3,J],C[1,J],O[2,J],O[3,J],O[4,J],O[5,J]
410 FOR I=1 TC 5
420C[I,J]=100*(O[I,J]-S[I,J])/S[I,J]
430 NEXT I
440 WRITE (15,460)O[1,J],O[2,J],O[3,J],O[4,J],O[5,J.]
450 FORNAT F10.0,5F10.4
460 FORLAT 10X,5E10.2
4 7 0 ~ P R I N T ~
4 8 0 ~ N E X T ~ J ~ J
4 9 0 ~ P R I N T ~
500 PRINT "Coefficient matrix:"
510 FGR I=1 TO 5
520 FiRITE (15,530) B [1,I],B[2,I],B[3,I],B[4,I],B[5,I]
530 FORMAT SEIO.3
540 NEXT I
550 ENE
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[^0]:    * The use of trade names is for illustrative purposes only and does not imply endorsement or recommendation by the U.S. Geological Survey.

